

No documents - You can use a calculator

Please report clear and detailed answers explicitly, as well as all the formulas you use and all the theorems you invoke.

**Problem 1 :**

1) Yes, the security  $i = 0$  is a numeraire because

$$\forall j \in [1, 4], \quad S_1^{(0)}(\omega_j) = 1.05 > 1.$$

The risk-free rate  $r$  is

$$(1 + r) \times 1 = 1.05 \Rightarrow r = 0.05 = 5\%.$$

2) The security

$$X = (11.05, 13.05, 15.05, 15.05)$$

is attainable if there exists  $\theta = (\theta_0, \theta_1, \theta_3) \in \mathbb{R}^3$  such that

$$X = \theta S_1 \Leftrightarrow S_1^T \theta^T = X^T$$

where

$$S_1 = \begin{pmatrix} 1.05 & 1.05 & 1.05 & 1.05 \\ 4 & 2 & 5 & 3 \\ 1 & 4 & 2 & 4 \end{pmatrix}.$$

We need to solve

$$\begin{pmatrix} 1.05 & 4 & 1 \\ 1.05 & 2 & 4 \\ 1.05 & 5 & 2 \\ 1.05 & 3 & 4 \end{pmatrix} \begin{pmatrix} \theta_0 \\ \theta_1 \\ \theta_2 \end{pmatrix} = \begin{pmatrix} 11.05 \\ 13.05 \\ 15.05 \\ 15.05 \end{pmatrix}$$

which we can write as

$$\begin{array}{ccc|c} 1.05 & 4 & 1 & 11.05 \\ 1.05 & 2 & 4 & 13.05 \\ 1.05 & 5 & 2 & 15.05 \\ 1.05 & 3 & 4 & 15.05 \end{array} \quad \begin{array}{l} L_2 \leftarrow L_2 - L_1 \\ L_3 \leftarrow L_3 - L_1 \\ L_4 \leftarrow L_4 - L_1 \end{array}$$

$$\begin{array}{l}
\begin{array}{ccc|c}
1.05 & 4 & 1 & 11.05 \\
0 & -2 & 3 & 2 \\
\Leftrightarrow & 0 & 1 & 4 \\
& 0 & -1 & 3 \\
& & & 4
\end{array} & \begin{array}{l} \\ \\ L_3 \leftarrow 2L_3 + L_2 \\ L_4 \leftarrow 2L_4 - L_2 \end{array} \\
\\
\begin{array}{ccc|c}
1.05 & 4 & 1 & 11.05 \\
0 & -2 & 3 & 2 \\
\Leftrightarrow & 0 & 0 & 5 \\
& & & 10 \\
& 0 & 0 & 3 \\
& & & 6
\end{array} & \begin{array}{l} \\ \\ L_3 \leftarrow L_3/5 \\ L_4 \leftarrow L_4/3 \end{array} \\
\\
\begin{array}{ccc|c}
1.05 & 4 & 1 & 11.05 \\
0 & -2 & 3 & 2 \\
\Leftrightarrow & 0 & 0 & 1 \\
& & & 2 \\
& 0 & 0 & 1 \\
& & & 2
\end{array} & \begin{array}{l} \\ \\ L_2 \leftarrow L_2 - 3L_3 \\ \\ \end{array} \\
\\
\begin{array}{ccc|c}
1.05 & 4 & 1 & 11.05 \\
0 & -2 & 0 & -4 \\
\Leftrightarrow & 0 & 0 & 1 \\
& & & 2 \\
& 0 & 0 & 1 \\
& & & 2
\end{array} & \begin{array}{l} \\ \\ L_2 \leftarrow L_2/(-2) \\ \\ \end{array} \\
\\
\begin{array}{ccc|c}
1.05 & 4 & 1 & 11.05 \\
0 & 1 & 0 & 2 \\
\Leftrightarrow & 0 & 0 & 1 \\
& & & 2 \\
& 0 & 0 & 1 \\
& & & 2
\end{array} & \begin{array}{l} \\ \\ L_1 \leftarrow L_1 - 4L_2 - L_3 \\ \\ \end{array} \\
\\
\begin{array}{ccc|c}
1.05 & 0 & 0 & 1.05 \\
0 & 1 & 0 & 2 \\
\Leftrightarrow & 0 & 0 & 1 \\
& & & 2 \\
& 0 & 0 & 1 \\
& & & 2
\end{array} & \begin{array}{l} \\ \\ L_1 \leftarrow L_1/1.05 \\ \\ \end{array} \\
\\
\begin{array}{ccc|c}
1 & 0 & 0 & 1 \\
0 & 1 & 0 & 2 \\
\Leftrightarrow & 0 & 0 & 1 \\
& & & 2 \\
& 0 & 0 & 1 \\
& & & 2
\end{array} & \begin{array}{l} \\ \\ \\ \\ \end{array}
\end{array}$$

We get the following equivalent portfolio

$$(\theta_0, \theta_1, \theta_2) = (1, 2, 2)$$

and then, the security  $X$  is attainable.

- 3)  $\mathcal{Q} = (q_1, q_2, q_3, q_4)$  is a risk-neutral probability distribution if
- $\forall j \in \llbracket 1, 4 \rrbracket, \quad q_j > 0,$
  - $q_1 + q_2 + q_3 + q_4 = 1,$

— deflated security prices are martingales. That is,

$$\begin{aligned} \forall i \in \llbracket 0, 2 \rrbracket : \quad \frac{S_0^{(i)}}{1} &= \mathbb{E}^{\mathcal{Q}} \left( \frac{S_1^{(i)}}{1+r} \right) = \frac{1}{1+r} \mathbb{E}^{\mathcal{Q}} \left( S_1^{(i)} \right) \\ &= \frac{1}{1+r} \sum_{j=1}^4 q_j S_1^{(i)}(\omega_j) \end{aligned}$$

that is equivalent to

$$S_0 = \frac{1}{1+r} S_1 \begin{pmatrix} q_1 \\ q_2 \\ q_3 \\ q_4 \end{pmatrix} \Leftrightarrow S_1 \begin{pmatrix} q_1 \\ q_2 \\ q_3 \\ q_4 \end{pmatrix} = 1.05 S_0.$$

We want solve

$$\begin{array}{cccc|cc} 1.05 & 1.05 & 1.05 & 1.05 & 1.05 & L_1 \leftarrow L_1/1.05 \\ 4 & 2 & 5 & 3 & 1.05 \times 3 & \\ 1 & 4 & 2 & 4 & 1.05 \times 3 & \\ \hline 1 & 1 & 1 & 1 & 1 & \\ \Leftrightarrow 4 & 2 & 5 & 3 & 3.15 & L_2 \leftarrow L_2 - 4L_1 \\ 1 & 4 & 2 & 4 & 3.15 & L_3 \leftarrow L_3 - L_1 \\ \hline 1 & 1 & 1 & 1 & 1 & \\ \Leftrightarrow 0 & -2 & 1 & -1 & -0.85 & L_2 \leftarrow L_2 + L_3 \\ 0 & 3 & 1 & 3 & 2.15 & \\ \hline 1 & 1 & 1 & 1 & 1 & \\ \Leftrightarrow 0 & 1 & 2 & 2 & 1.3 & \\ 0 & 3 & 1 & 3 & 2.15 & L_3 \leftarrow L_3 - 3L_2 \\ \hline 1 & 1 & 1 & 1 & 1 & \\ \Leftrightarrow 0 & 1 & 2 & 2 & 1.3 & \\ 0 & 0 & -5 & -3 & -1.75 & L_3 \leftarrow L_3/(-5) \\ \hline 1 & 1 & 1 & 1 & 1 & \\ \Leftrightarrow 0 & 1 & 2 & 2 & 1.3 & L_2 \leftarrow L_2 - 2L_3 \\ 0 & 0 & 1 & 3/5 & 0.35 & \\ \hline 1 & 1 & 1 & 1 & 1 & L_1 \leftarrow L_1 - L_2 - L_3 \\ \Leftrightarrow 0 & 1 & 0 & 4/5 & 0.6 & \\ 0 & 0 & 1 & 3/5 & 0.35 & \\ \hline 1 & 0 & 0 & -2/5 & 0.05 & L_1 \leftarrow L_1 - L_2 - L_3 \\ \Leftrightarrow 0 & 1 & 0 & 4/5 & 0.6 & \\ 0 & 0 & 1 & 3/5 & 0.35 & \end{array}$$

Then, for all  $q_4 \in \mathbb{R}$ ,

$$\left( 0.05 + \frac{2}{5}q_4, 0.6 - \frac{4}{5}q_4, 0.35 - \frac{3}{5}q_4, q_4 \right)$$

is solution of

$$S_0 = \frac{1}{1+r} S_1 \begin{pmatrix} q_1 \\ q_2 \\ q_3 \\ q_4 \end{pmatrix}.$$

Since, we want  $q_1 > 0$ ,  $q_2 > 0$ ,  $q_3 > 0$  and  $q_4 > 0$ , we need

$$\begin{cases} 0.05 + \frac{2}{5}q_4 > 0 \\ 0.6 - \frac{4}{5}q_4 > 0 \\ 0.35 - \frac{3}{5}q_4 > 0 \\ q_4 > 0 \end{cases} \Leftrightarrow \begin{cases} q_4 > -\frac{5 \times 0.05}{2} = 0.125 \\ q_4 < \frac{5 \times 0.6}{4} = 0.75 \\ q_4 < \frac{5 \times 0.35}{3} \approx 0.583 \\ q_4 > 0 \end{cases}$$

Then, for all  $q_4 \in ]0.125, 0.583[$ ,

$$\left( 0.05 + \frac{2}{5}q_4, 0.6 - \frac{4}{5}q_4, 0.35 - \frac{3}{5}q_4, q_4 \right)$$

is a risk-neutral probability distribution. For example, with  $q_4 = 0.2$ , we obtain that

$$\mathcal{Q} = (0.13, 0.44, 0.23, 0.20)$$

is a risk-neutral probability distribution.

- 4) Since there exists a risk-neutral probability distribution, the fundamental theorem of asset pricing (part 1) gives us that there is no possible arbitrage. Then, there is no arbitrage of type A and then, by a proposition of the course, the linear pricing hypothesis is satisfied.
- 5) By the fundamental theorem of asset pricing, we know that there exists a vector of states prices. Moreover, since the linear pricing hypothesis is satisfied,  $\pi = (\pi_1, \pi_2, \pi_3, \pi_4)$  is a vector of state prices if

$$\begin{cases} (\pi_1, \pi_2, \pi_3, \pi_4) > 0, \\ S_1 \begin{pmatrix} \pi_1 \\ \pi_2 \\ \pi_3 \\ \pi_4 \end{pmatrix} = S_0. \end{cases}$$

We saw in question 3) that a risk-neutral probability distribution if and only

if

$$S_0 = \frac{1}{1+r} S_1 \begin{pmatrix} q_1 \\ q_2 \\ q_3 \\ q_4 \end{pmatrix} \quad \text{with} \quad \begin{pmatrix} q_1 \\ q_2 \\ q_3 \\ q_4 \end{pmatrix} > 0.$$

Then, we obtain

$$\begin{aligned} \pi &= (\pi_1, \pi_2, \pi_3, \pi_4) \text{ is a vector of state prices} \\ \Leftrightarrow \frac{\pi}{1+r} &= \left( \frac{\pi_1}{1+1}, \frac{\pi_2}{1+r}, \frac{\pi_3}{1+r}, \frac{\pi_4}{1+r} \right) \text{ is a risk-neutral probability} \\ &\hspace{15em} \text{distribution.} \end{aligned}$$

Then, for all  $q_4 \in ]0.125, 0.583[$ ,

$$(1+r) \times \left( 0.05 + \frac{2}{5}q_4, 0.6 - \frac{4}{5}q_4, 0.35 - \frac{3}{5}q_4, q_4 \right)$$

is a vector of state prices. For example, with  $q_4 = 0.2$ , we obtain that

$$\pi = (0.1365, 0.4620, 0.2415, 0.2100)$$

is a vector of state prices.

- 6) Since linear pricing hypothesis holds, the fair price at time  $t = 0$  associated to portfolio  $\theta = (1, 2, 2)$  is

$$\begin{aligned} P &= \theta S_0 = (1, 2, 2) \begin{pmatrix} 1 \\ 3 \\ 3 \end{pmatrix} \\ &= 1 + 2 \times 3 + 2 \times 3 = 1 + 6 + 6 = 13. \end{aligned}$$

Then, it is the fair price.

### Problem 2 :

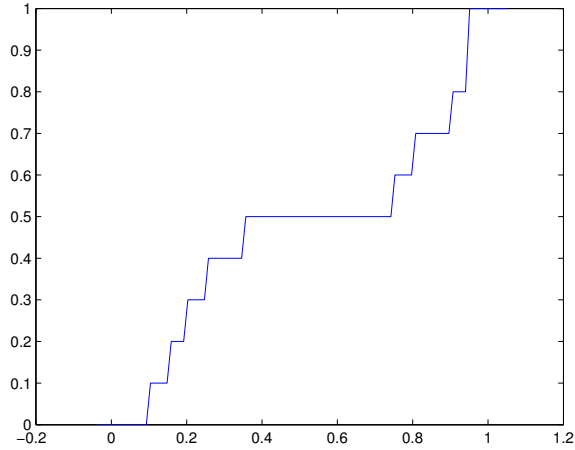
- 1) The cumulative function is defined by

$$F(x) = \int_{-\infty}^x f(t) dt$$

where  $f$  is the density function. If we have a random generator  $(y_i)_{1 \leq i \leq N}$ . We can estimate the cumulative function associate to  $(y_i)_{1 \leq i \leq N}$  with

$$F_N(x) = \frac{\#\{i : y_i \leq x\}}{N}.$$

- 2) We obtain



3) The density function for a uniform law is

$$f(x) = \begin{cases} 1 & \text{if } 0 \leq x \leq 1, \\ 0 & \text{elsewhere.} \end{cases}$$

Then, the cumulative function for a uniform law is

$$F(x) = \int_{-\infty}^x f(t)dt = \begin{cases} \int_{-\infty}^x 0dt & \text{if } x \leq 0, \\ \int_{-\infty}^0 0dt + \int_0^x 1dt & \text{if } 0 \leq x \leq 1, \\ \int_{-\infty}^0 0dt + \int_0^1 1dt + \int_1^x 0dt & \text{elsewhere.} \end{cases}$$

$$= \begin{cases} 0 & \text{if } x \leq 0, \\ x & \text{if } 0 \leq x \leq 1, \\ 1 & \text{elsewhere.} \end{cases}$$

4) This is not a uniform distribution since the cumulative function is far from the line of equation  $y = x$ .

**Problem 3 :**

1) We have  $y_0 = 107$ ,  $a = 327$ ,  $c = 1$ ,  $m = 1000$  and linear congruencies sequence is defined by

$$y_{n+1} = a \times y_n + c \pmod{m}.$$

Then, we obtain

$$\begin{aligned} y_1 &= a \times y_0 + c \pmod{m} \\ &= 327 \times 107 + 1 \pmod{1000} = 34990 \pmod{1000} \\ &= 990. \end{aligned}$$

Similarly, we obtain

$$\begin{aligned}
 y_2 &= 327 \times 990 + 1 \pmod{1000} \\
 &= 323731 \pmod{1000} = 731 \\
 y_3 &= 327 \times 731 + 1 \pmod{1000} \\
 &= 239038 \pmod{1000} = 38 \\
 y_4 &= 327 \times 38 + 1 \pmod{1000} \\
 &= 12427 \pmod{1000} = 427.
 \end{aligned}$$

2) We just have to divide by  $m = 1000$ , we obtain

$$z_0 = 0.107, \quad z_1 = 0.99, \quad z_2 = 0.731, \quad z_3 = 0.038, \quad z_4 = 0.427.$$

3)  $f$  is a density function because  $f$  is positive and

$$\int_{-\infty}^{+\infty} f(x)dx = \int_0^{3^{1/3}} x^2 dx = \left[ \frac{x^3}{3} \right]_0^{3^{1/3}} = \frac{1}{3} \left( 3^{1/3} \right)^3 = \frac{1}{3} 3 = 1.$$

4) The cumulative function  $F$  of the random variable  $X$  is

$$\begin{aligned}
 F(x) &= \int_{-\infty}^x f(t)dt = \begin{cases} \int_{-\infty}^x 0dt & \text{if } x \leq 0, \\ \int_{-\infty}^0 0dt + \int_0^x t^2 dt & \text{if } 0 \leq x \leq 3^{1/3}, \\ \int_{-\infty}^0 0dt + \int_0^{3^{1/3}} t^2 dt + \int_{3^{1/3}}^x 0dt & \text{elsewhere} \end{cases} \\
 &= \begin{cases} 0 & \text{if } \leq 0, \\ \left[ \frac{x^3}{3} \right]_0^x & \text{if } 0 \leq x \leq 3^{1/3}, \\ \left[ \frac{x^3}{3} \right]_0^{3^{1/3}} & \text{elsewhere} \end{cases} = \begin{cases} 0 & \text{if } \leq 0, \\ \frac{x^3}{3} & \text{if } 0 \leq x \leq 3^{1/3}, \\ 1 & \text{elsewhere.} \end{cases}
 \end{aligned}$$

5) We know that if  $Z$  is a uniform random variable on  $[0, 1]$ , the random variable  $X = F^{-1}(Z)$  has density  $f$ . Then,

$$F^{-1}(z_1), \quad F^{-1}(z_2), \quad F^{-1}(z_3)$$

are three realizations of random variable  $X$ . We just have to prove that  $F$  is bijective and to compute  $F^{-1}$ . With the question 4), we obtain that  $F$  is a bijective function from  $[0, 3^{1/3}]$  to  $[0, 1]$ . We need to compute  $F^{-1}(y)$  for some  $y \in [0, 1]$ , we have

$$\begin{aligned}
 y = F(x) &\Leftrightarrow y = \frac{x^3}{3} \Leftrightarrow x^3 = 3y \Leftrightarrow x = (3y)^{1/3} \\
 &\Leftrightarrow x = (3y)^{1/3} = F^{-1}(y).
 \end{aligned}$$

Then,

$$(3z_1)^{1/3} = 1.44, \quad (3z_2)^{1/3} = 1.299, \quad (3z_3)^{1/3} = 0.485$$

are three realizations of random variable  $X$ .

6) We have

```
function z= my_generator(a,c,m,y0,k)
    y=congruencial(a,c,m,y0,k);
    for i=1:1:k
        z(i)=(3*y(i))^(1/3)
    end
end
```

7) We know that if  $Z_1$  and  $Z_2$  are two uniform random variables on  $[0, 1]$ , the random variable

$$\sigma \sqrt{-2 \ln(Z_1)} \cos(2 \pi Z_2) + \mu,$$

is a normal random variable with mean  $\mu$  and standard deviation  $\sigma$ . In our case, we want a standard normal distribution, then we use  $\mu = 0$  and  $\sigma = 1$  in the Box and Miller formula. Then,

$$\begin{aligned} \sqrt{-2 \ln(z_1)} \cos(2 \pi z_2) &= -0.0169, \\ \sqrt{-2 \ln(z_3)} \cos(2 \pi z_4) &= -2.2931 \end{aligned}$$

are two realizations of a standard normal random variable.

#### Problem 4 :

We assume that the following data are values of the price of an security at time  $T = 100$  obtained by performing a Monte-Carlo simulations :

$$125, 135, 95, 100.$$

Assume that this security follows a binomial model with  $T = 100$ .

1) The four simulations give us the following payoffs :

$$15, 25, 0, 0.$$

The average of these four values is

$$\theta_4 = \frac{15 + 25 + 0 + 0}{4} = \frac{40}{4} = 10.$$

The value of the call obtained with the Monte-Carlo method is

$$C = \frac{\theta_4}{(1+r)^{100}} = \frac{10}{(1+0.02)^{100}} \approx 1.38.$$

2) To give a confidence interval, we firstly need to compute the standard devia-



tion  $s$  of the four previous value, we have

$$\begin{aligned} s &= \sqrt{\frac{(15 - 10)^2 + (25 - 10)^2 + (0 - 10)^2 + (0 - 10)^2}{4}} \\ &= \sqrt{\frac{5^2 + 15^2 + 10^2 + 10^2}{4}} \approx 10.60. \end{aligned}$$

Since the confidence level is 95%, we have  $Z = 1.96$  and the confidence interval is  $[l, u]$  with

$$\begin{aligned} l &= \theta_4 - Z \frac{s}{\sqrt{n}} = 10 - 1.96 \times \frac{10.60}{\sqrt{4}} \approx -0.39 \\ u &= \theta_4 + Z \frac{s}{\sqrt{n}} = 10 + 1.96 \times \frac{10.60}{\sqrt{4}} \approx 20.39, \end{aligned}$$

because

$$\begin{aligned} P\left(-Z \leq \frac{\theta_n - \theta}{s/\sqrt{n}} \leq Z\right) &= 95\% = 0.95 \\ \Leftrightarrow P\left(\theta_n - Z \times \frac{s}{\sqrt{n}} \leq \theta \leq \theta_n + Z \times \frac{s}{\sqrt{n}}\right) &= 0.95 \end{aligned}$$

and it is assumed with the Central Limit Theorem that  $\frac{\theta_n - \theta}{s/\sqrt{n}}$  is a standard normal distribution.

We assume that the value obtained in Question 1) are used as the first step of a two phases methods for performing a confidence interval.

3) We have

$$\begin{aligned} P\left(-Z \leq \frac{\theta_n - \theta}{s\sqrt{n}} \leq Z\right) &= 95\% = 0.95 \\ \Leftrightarrow P\left(-Z \times \frac{s}{\sqrt{n}} \leq \theta_n - \theta \leq Z \times \frac{s}{\sqrt{n}}\right) &= 0.95 \\ \Leftrightarrow P\left(|\theta_n - \theta| \leq Z \times \frac{s}{\sqrt{n}}\right) &= 0.95. \end{aligned}$$

Then, the absolute error  $|\theta_n - \theta|$  is less than  $\epsilon = 3$  if

$$Z \times \frac{s}{\sqrt{n}} \leq \epsilon \Leftrightarrow n \geq \frac{s^2 \times Z^2}{\epsilon^2} \approx \frac{10.60^2 \times 1.96^2}{3^2} = 47.96.$$

Then, to obtained an absolute error of at most 3 with a confidence level of 95%, we need at least

$$n = 48$$

simulations.

4) We have

$$\begin{aligned} P\left(-Z \leq \frac{\theta_n - \theta}{s/\sqrt{n}} \leq Z\right) &= 95\% = 0.95 \\ \Leftrightarrow P\left(-Z \times \frac{s}{\sqrt{n}} \leq \theta_n - \theta \leq Z \times \frac{s}{\sqrt{n}}\right) &= 0.95 \\ \Leftrightarrow P\left(|\theta_n - \theta| \leq Z \times \frac{s}{\sqrt{n}}\right) &= 0.95 \\ \Leftrightarrow P\left(\left|\frac{\theta_n - \theta}{\theta}\right| \leq \frac{Z \times s}{|\theta|\sqrt{n}}\right) &= 0.95. \end{aligned}$$

Then, the relative error  $\left|\frac{\theta_n - \theta}{\theta}\right|$  is less than  $\epsilon = 3\% = 0.03$  if

$$\frac{Z \times s}{|\theta_4|\sqrt{n}} \leq \epsilon \Leftrightarrow n \geq \frac{s^2 \times Z^2}{\theta_4^2 \times \epsilon^2} \approx \frac{10.60^2 \times 1.96^2}{10^2 \times 0.03^2} = 4796.02.$$

Then, to obtain a relative error of at most  $3\% = 0.03$  with a confidence level of  $95\%$ , we need at least

$$n = 4797$$

simulations.